# RANDOM PROCESSES. THE FINAL TEST.

Dr. P. Chigansky 9:00-12:00, 20th, September, 2004

Student	$ID \cdot$	
Student	1D:	

- \* any supplementary material is allowed
- \* duration of the exam is 3 hours
- \* write <u>briefly</u> the main idea of your answers in the exam itself. If required, give the reference to your copybook, where you may place other technical details
- \* note that the problems are <u>not</u> in any monotonic order of complexity
- \* the total score of the exam is 120 points.
- \* good luck!

### 1

Problem 1.

Let X be a standard Gaussian random variable (a "message") to be transmitted over noisy channel, so that the following observation sequence is available to the receiver:

$$Y_n = a_n + b_n X + \xi_n, \quad n \ge 1$$

where  $\xi = (\xi_n)_{n\geq 1}$  is a standard Gaussian i.i.d. random sequence, independent of X and  $(a_n, b_n)_{n\geq 1}$  are real numbers, chosen by the transmitter and known also to the receiver.

(a) Find recursive formulae for the optimal receiver  $\widehat{X}_n = E(X|Y_1^n)$  and the corresponding mean square error  $P_n = E(X - \widehat{X}_n)^2$ .

(reference page \_\_\_\_)

(b) Let  $\gamma_n = E(a_n + b_n X)^2$  be the receiver output power. The optimal transmitter, which minimizes  $P_n$  and satisfies the power constraint  $\gamma_n \leq \gamma$  for any  $n \geq 1$  is

(1) 
$$a_n = 0, b_n = 1 \text{ and } P_n = \gamma$$

(2) 
$$a_n = 0, b_n = \sqrt{\gamma} \text{ and } P_n = 1/(1 + \gamma n)$$

(3) 
$$a_n = 0$$
,  $b_n = \sqrt{\gamma}$  and  $P_n = \gamma/(\gamma + n)$ 

(4) 
$$a_n = -\gamma$$
,  $b_n = \sqrt{\gamma}$  and  $P_n = 1/(1+n)$ 

(c) If $\xi_1$	were a non	${\it Gaussian}$	${\rm random}$	variable	with	zero	mean	and	unit
variance,	the optimal	transmitt	er/receiv	er pair n	night	attaiı	n		

- (1) smaller error than in the Gaussian case
- (2) larger error than in the Gaussian case

(reference page \_\_\_\_)

(d) Assume that the transmitter gets noiseless feedback from the receiver, so that only the coefficient  $a_n$  (and not  $b_n$ ) is allowed to depend on  $\{Y_1, ..., Y_{n-1}\}$ , the information passed to the receiver before n:

$$Y_n = a_n(Y_1^{n-1}) + b_n X + \xi_n.$$

Derive the equations for  $\widehat{X}_n = E(X|Y_1^n)$  and  $P_n = E(X-\widehat{X}_n)^2$ , the optimal receiver in this case.

- (e) Is  $Y = (Y_n)_{n \ge 1}$  a Gaussian process in (d)?
  - (1) Yes.
  - (2) No

Explain:

(f) The optimal transmitter, which minimizes  $P_n$  subject to the power constraint  $\gamma_n = E(a_n(Y_1^{n-1}) + b_n X)^2 \leq \gamma$  is

(1) 
$$a_n = 0, b_n = \sqrt{\gamma} \text{ and } P_n = 1/(1 + \gamma n)$$

(2) 
$$a_n = 0$$
,  $b_n = \sqrt{\gamma}$  and  $P_n = \gamma/(\gamma + n)$ 

(3) 
$$a_n = -b_n \hat{X}_{n-1}$$
,  $b_n = \sqrt{\gamma(1+\gamma)^{n-1}}$  and  $P_n = 1/(1+\gamma)^n$ 

(4) 
$$a_n = -b_n \hat{X}_{n-1}, b_n = \sqrt{\gamma^{n-1}(1+\gamma)} \text{ and } P_n = 1/(1+\gamma)^n$$

Hint: Convince yourself that

$$E(a_n(Y_1^{n-1}) + b_n X)^2 = E(a_n(Y_1^{n-1}) + b_n \widehat{X}_{n-1})^2 + b_n^2 P_{n-1}$$
we it with the equation for  $P_n$  (without explicitly solving it)

and use it with the equation for  $P_n$  (without explicitly solving it).

(g) Can the filtering error be improved, if  $b_n$  is also allowed to depend on  $\{Y_1,...,Y_{n-1}\}$ 

(1) Yes, by choosing

 $a_n = \underline{\hspace{1cm}}$  and  $b_n = \underline{\hspace{1cm}}$  which gives

 $P_n = \underline{\hspace{1cm}}$ 

(2) No.

Explain:

**Hint:** If  $\zeta_n$  are positive random variables then (why?)

$$E\prod_{\ell=1}^{n} \frac{1}{\zeta_{\ell}} \ge \exp\left\{-\sum_{\ell=1}^{n} \log E\zeta_{\ell}\right\}$$

 $({\rm reference\ page\ }\underline{\hspace{1cm}})$ 

# Problem 2.

Let  $X=(X_n)_{n\geq 1}$  be a sequence of i.i.d. random variables. For a fixed  $n\geq 1$  let  $Y^n$  be the vector with entries

$$Y^{n}(i) = X_{i} / \sqrt{X_{1}^{2} + ... + X_{n}^{2}}, \quad i = 1, ..., n$$

- (a) Assume  $EX_1^2=1$  and  $E|X_1|^p<\infty$  for any  $p\geq 1$ . Does the random sequence  $\sqrt{n}Y^n(1)$  converge as  $n\to\infty$ ?
  - (1) Yes,
    - $\square$  *P*-a.s.
    - $\square$  in probability
    - $\square$  in  $\mathbb{L}^2$
    - $\square$  in law

the limit is \_\_\_\_\_

(2) No.

Hint: use the law of large numbers

(b) A random vector Z in  $\mathbb{R}^n$  is said to have uniform distribution on the unit sphere in  $\mathbb{R}^n$ , if its Euclidian norm is unity and it's distribution is invariant under rotations, i.e for any orthogonal matrix U, such that  $U^{-1} = U^*$ , Z and UZ have the same distribution.

 $Y^n$  has uniform distribution on the unit sphere in  $\mathbb{R}^n$  for any n>1 if

- (1)  $X_1$  is Gaussian with zero mean
- (2)  $X_1$  is Bernulli, i.e.  $P(X_1 = \pm 1) = 1/2$
- (3)  $X_1$  takes values in  $\{\pm 1, \pm 2, ...\}$  and  $P(X_1 = \ell) = P(X_1 = -\ell)$

- (c) It is known that the uniform distribution on the unit sphere in  $\mathbb{R}^n$ , n > 1 is unique, i.e. there is only one distribution which is invariant under rotations. Let  $\mathbb{Z}^n$  be a random vector with this distribution. Then
  - (1)  $\sqrt{n}Z^n(1)$  converges weakly to a uniform random variable on [-1,1]
  - (2)  $\sqrt{n}Z^n(1)$  converges P-a.s. to a uniform random variable on [-1,1]
  - (3)  $\sqrt{n}Z^n(1)$  converges weakly to a standard Gaussian random variable
- (4)  $\sqrt{n}Z^n(1)$  converges P-a.s. to a standard Gaussian random variable Explain:

# Problem 3.

Let  $X = (X_n)_{n \ge 0}$  be a binary Markov chain , switching between 0 and 1 with transition probabilities

$$\lambda_0 = P(X_n = 0 | X_{n-1} = 0), \quad \lambda_1 = P(X_n = 1 | X_{n-1} = 1)$$

and equiprobable initial distribution. The observation process is given by

$$Y_n = X_n + \varepsilon_n, \quad n \ge 1$$

where  $\varepsilon = (\varepsilon_n)_{n\geq 1}$  is an i.i.d. sequence, independent of X and  $\varepsilon_1$  has probability density f(x).

Introduce the process  $Z = (Z_n)_{n \ge 0}$ 

$$Z_n = \prod_{k=0}^n X_k.$$

- (a) Does  $Z_n$  converge?
  - (1) Yes,

 $\square$  *P*-a.s.

 $\square$  in probability

 $\square$  in  $\mathbb{L}^2$ 

 $\square$  in law

the limit is \_\_\_\_\_

(2) No.

Explain:

<b>(b)</b> Does the sequence $\widehat{Z}_n$	$= P(Z_n = 1 Y_1^n) \text{ converge } ?$
(1) Yes,	
☐ in probability	

$$\Box$$
 in  $\mathbb{L}^2$ 
 $\Box$  in law
the limit is \_\_\_\_\_

(2) No.

Explain:

- (c) Is Z a Markov process ?
  - (1) Yes, the transition probabilities are

$$P(Z_n = 0|Z_{n-1} = 0) =$$
\_\_\_\_\_\_  
 $P(Z_n = 1|Z_{n-1} = 1) =$ \_\_\_\_\_

(2) No.

Explain:

(d) The conditional probability  $\pi_n = P(X_n = 1 | Y_1^n)$  satisfies the recursion:

$$\pi_n =$$

 $({\rm reference\ page\ } \underline{\hspace{1cm}})$ 

(e) The conditional probability  $\widehat{Z}_n = P(Z_n = 1|Y_1^n)$  satisfies

$$1) \quad \widehat{Z}_n = \widehat{Z}_{n-1} \pi_n$$

$$2) \quad \widehat{Z}_n = \frac{\widehat{Z}_{n-1}}{\pi_{n-1}} \pi_n$$

3) 
$$\widehat{Z}_n = \frac{\lambda_1 \widehat{Z}_{n-1} + (1 - \lambda_0)(1 - \widehat{Z}_{n-1})}{\lambda_1 \pi_{n-1} + (1 - \lambda_0)(1 - \pi_{n-1})} \pi_n$$

4) 
$$\widehat{Z}_n = \frac{\lambda_1(1-\widehat{Z}_{n-1}) + (1-\lambda_0)\widehat{Z}_{n-1}}{\lambda_1(1-\pi_{n-1}) + (1-\lambda_0)\pi_{n-1}} \pi_n$$

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### Problem 4.

Let  $B = (B_t)_{t\geq 0}$  be a Wiener process. It turns out that any random variable X with  $EX^2 < \infty$ , generated by a trajectory of B on the interval [0,1], obeys the representation via Itô integral:

$$X = EX + \int_0^1 Z_s dB_s$$

for some random process  $Z = (Z_t)_{0 \le t \le 1}$ . Use the Itô formula to find  $Z_t$  for each of the following random variables:

(a) 
$$B_1 = \underline{\hspace{1cm}} + \int_0^1 \underline{\hspace{1cm}} dB_t$$

(reference page \_\_\_\_)

(b) 
$$B_1^2 = \underline{\hspace{1cm}} + \int_0^1 \underline{\hspace{1cm}} dB_t$$

<sup>&</sup>lt;sup>1</sup>or more precisely X is measurable w.r.t.  $\mathcal{F}_1^B = \sigma\{B_s, 0 \le s \le 1\}$ 

(c) 
$$\int_0^1 B_s ds = \underline{\qquad} + \int_0^1 \underline{\qquad} dB_t$$

**Hint:** Apply the Itô formula to  $tB_t$ 

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(d) 
$$B_1^3 = \underline{\hspace{1cm}} + \int_0^1 \underline{\hspace{1cm}} dB_t$$

 $({\rm reference\ page\ } \underline{\hspace{1cm}})$ 

(e) 
$$\exp(B_1) = \underline{\hspace{1cm}} + \int_0^1 \underline{\hspace{1cm}} dB_t$$

**Hint:** Apply the Itô formula to  $\exp\{B_t - t/2\}$ .

$$\sin(B_1) = \underline{\qquad} + \int_0^1 \underline{\qquad} dB_t$$

**Hint:** Apply the Itô formula to  $\exp\{t/2\}\sin(B_t)$ .

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